

### Extensions to linear and logistic regressions

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#### Conditional logistic regression

- When?
- What?
- How?

#### Other methods for analyzing binary data

Models for **relative risks**  
Models for **risk differences**

#### Clustered data / data with several random components

**Continuous outcome**  
**Dichotomous outcome**

#### Clustered binary data with one random components

#### Nonlinear regression models

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### Conditional logistic regression When

Used in two situations:

1. **Matched** studies (binary response).
2. **Unmatched** studies with a **confounder** with many **distinct values**.

In 1. the models correspond to **the way data was collected**.

In 2. the method adjust for a '**mathematical**' flaw in the unconditional method.

An example of situation 2. the confounder is "*kommune*" having 275 distinct values.

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### Conditional logistic regression What

The logistic regression model (outcome disease yes/no):

$$\ln(\text{odds}) = \alpha + \sum_{i=1}^k (\beta_i \cdot x_i)$$

$\ln(\text{odds})$  in reference

$\ln(\text{odds ratios})$

Suppose the model above hold in each strata:

$$\ln(\text{odds}) = \alpha_s + \sum_{i=1}^k (\beta_i \cdot x_i)$$

$\ln(\text{odds})$  in reference

**different in each strata**

$\ln(\text{odds ratios})$

**the same in each strata**

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### Conditional logistic regression What

$$\ln(\text{odds}) = \alpha_s + \sum_{i=1}^k (\beta_i \cdot x_i)$$

$\ln(\text{odds})$  different in each strata

**We are not interested in these !**

In a **matched** study these are 'controlled'.

In a **conditional** logistic regression one '**condition on the odds in each strata**', i.e. these case/control ratio.

In the conditional model the  $\alpha$ 's **disappear** !

The  $\beta$ 's, the log OR's, are still in and **can be estimated**.

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### Conditional logistic regression How

**It is easy !**

You need a statistical software package.

A package made for **research in epidemiology**

Not in social science

**Not SPSS**

But **STATA**, **EPICURE**, **EPILOG**, **EGRET**,  
**EPIINFO(2000)** and **SAS** can do it.

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### Conditional logistic regression How

An example using **STATA**

A study of cancer in the oral cavity

Matched on **gender** and **10 years age groups**

Ten strata (*genage*)

Here we focus on

*textile-worker* and

*life time consumption of alcohol* (three groups)

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Conditional logistic regression  
How

logistic regression in STATA

xi:logit cancer textile i.alkcon i.genage

Part of the output:

	cancer	Coef.	Std. Err.	z	P> z	CI
textile		.5022	.4141	1.213	0.225	-.3094 1.3139
_Ialkcon_1		.4628	.2823	1.639	0.101	-.0905 1.0163
_Ialkcon_2		2.7165	.3232	8.404	0.000	2.0829 3.3501
_Igenage_2		.2450	1.2514	0.196	0.845	-2.2075 2.6977
_Igenage_3		-.4940	.5503	-0.898	0.369	-1.5726 .5846
_Igenage_4		.1798	.6406	0.281	0.779	-1.0758 1.4353
_Igenage_5		-.2895	.5482	-0.529	0.597	-1.3644 .7844
_Igenage_6		.2127	.6268	0.340	0.734	-1.0147 1.4401
_Igenage_7		-.2305	.5355	-0.431	0.667	-1.2802 .8190
_Igenage_8		.5507	.5803	1.046	0.295	-.4809 1.5825
_Igenage_9		.0315	.5884	0.054	0.957	-1.1217 1.1847
_Igenage_10		-.5572	.5595	-0.996	0.319	-1.5954 1.6539
_const		-1.4692	.4762	-3.085	0.002	-2.4027 -.5356

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Conditional logistic regression in STATA

The syntax:

xi:clogit cancer textile i.alkcon, group(genage)

Part of the output:

	cancer	Coef.	Std. Err.	z	P> z	CI
textile		.4929	.4103	1.201	0.230	-.3112 1.2971
_Ialkcon_1		.452	.27923	1.621	0.105	-.094 .9999
_Ialkcon_2		2.660	.31936	8.332	0.000	2.034 3.2868

xi:clogit cancer textile i.alkcon, group(genage) or

	cases	Odds Ratio	Std. Err.	z	P> z	[95% Conf. Interval]
textile		1.63708	.6717022	1.20	0.230	.732517 3.658661
_Ialkcon_1		1.572508	.4390957	1.62	0.105	.909724 2.718168
_Ialkcon_2		14.30908	4.569879	8.33	0.000	7.651811 26.75835

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Other methods to analysis of binary response data  
Relative Risk models

Logistic regression model focus on the Odds Ratios

This is the correct thing to do in case-control studies.

In follow-up studies Relative Risk is often the appropriate measure of association, (personal risk).

I.e. a model like this might be more relevant:

$$\Pr(\text{event}) = p_0 \times RR_1 \times RR_2 \times RR_3$$
$$\ln\{\Pr(\text{event})\} = \ln(p_0) + \ln(RR_1) + \ln(RR_2) + \ln(RR_3)$$
$$\ln\{\Pr(\text{event given the covariates})\} = \alpha + \sum_{i=1}^p (\beta_i \cdot x_i)$$

That is linear on log-probability scale

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Other methods to analysis of binary response data  
Relative Risk models

$$\ln\{\Pr(\text{event given the covariates})\} = \alpha + \sum_{i=1}^p (\beta_i \cdot x_i)$$

Such a model modelling the relative risk can easily be fitted by many programs (not SPSS).

Logistic regression in STATA:

xi: logit obese age i.sex

or

xi: glm obese age i.sex, fam(bin) link(logit)

Relative risk model:

xi: glm obese age i.sex, fam(bin) link(log)

The link is log instead of logit

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Other methods to analysis of binary response data  
Risk difference models

Logistic regression model focus on the Odds Ratios

This is the correct thing to do in case-control studies.

In follow-up studies Risk Difference is often the appropriate measure of association, (community effect).

I.e. a model like this might be more relevant:

$$\Pr(\text{event}) = p_0 + RD_1 + RD_2 + RD_3$$
$$\Pr(\text{event given the covariates}) = \alpha + \sum_{i=1}^p (\beta_i \cdot x_i)$$

That is linear on probability scale

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Other methods to analysis of binary response data  
Risk difference models

$$\Pr(\text{event given the covariates}) = \alpha + \sum_{i=1}^p (\beta_i \cdot x_i)$$

Such a model modelling the risk difference can easily be fitted by many programs (not SPSS).

Logistic regression in STATA:

xi: logit obese age i.sex

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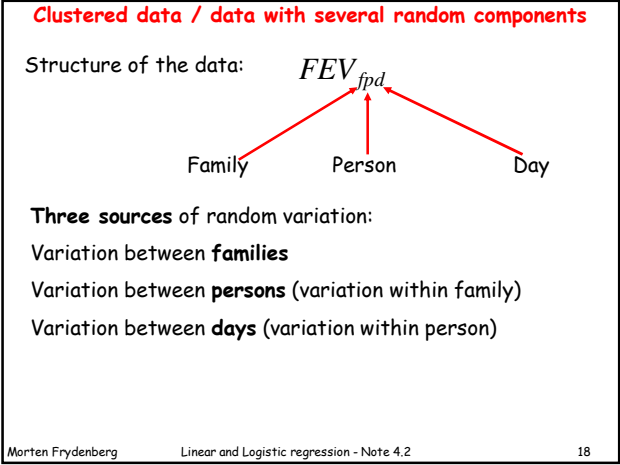
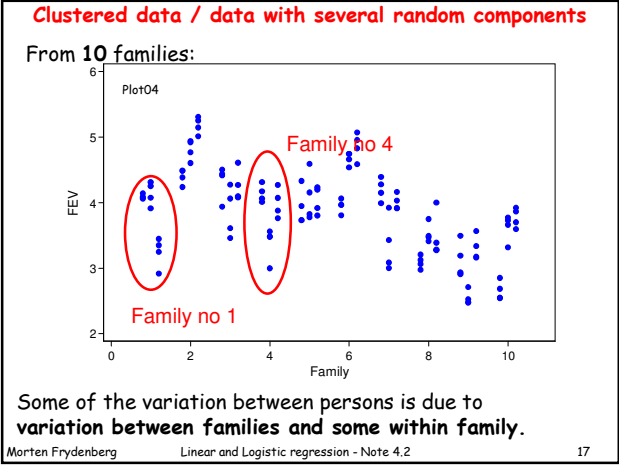
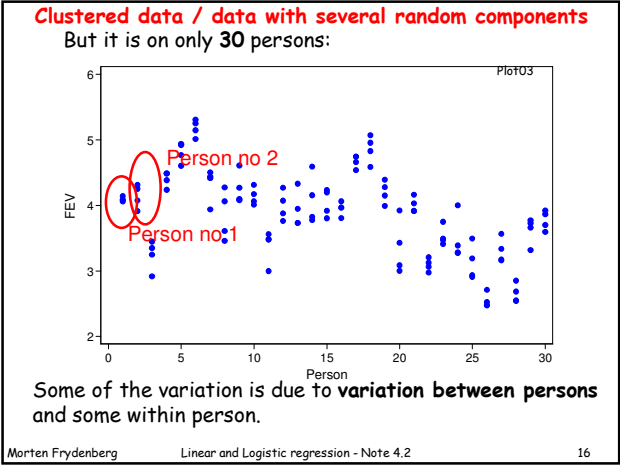
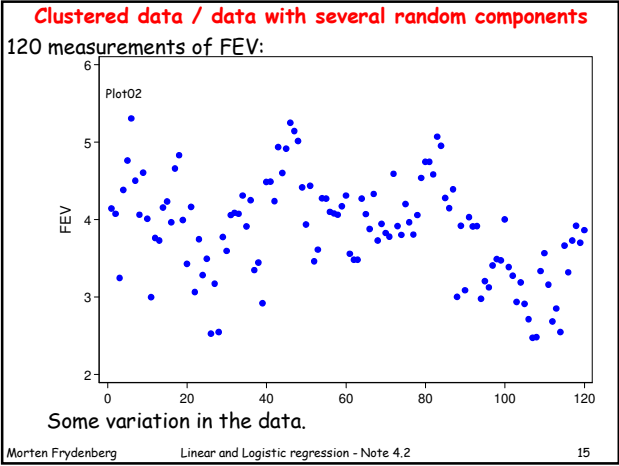
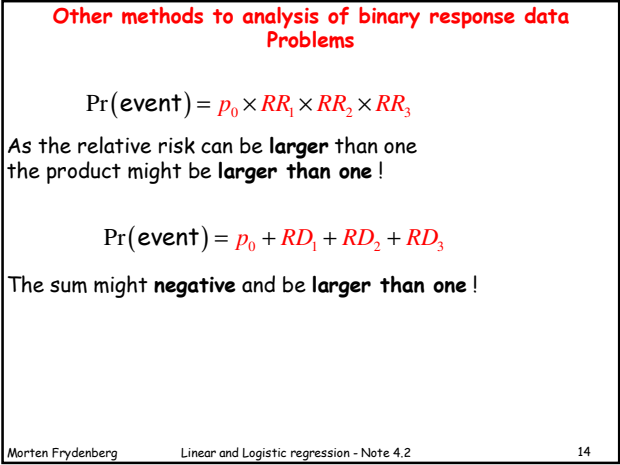
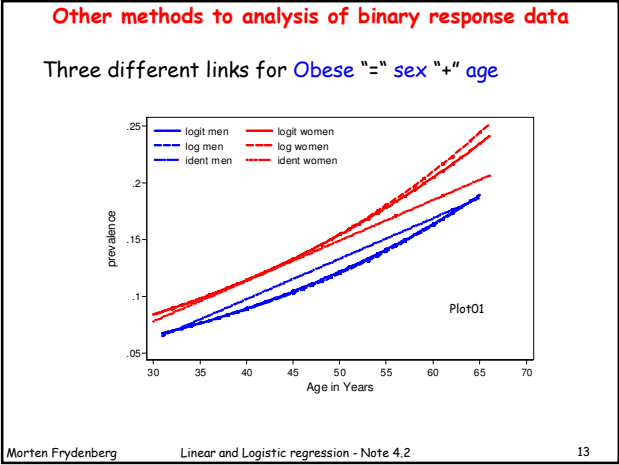
xi: glm obese age i.sex, fam(bin) link(logit)

Risk difference model:

xi: glm obese age i.sex, fam(bin) link(id)

The link is identity instead of logit

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Clustered data / data with several random components

Factors of interest:

household Income

Constant within family

Urbanization

Constant within family

Age

Constant within person; varies within family

Sex

Constant within person; varies within family

Grass pollen

Constant within day; varies within person

A model:

$$FEV = \beta_0 + \beta_I \cdot I + \beta_U \cdot U + \beta_A \cdot A + \beta_S \cdot S + \beta_G \cdot G$$

+random variation

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Clustered data / data with several random components

$$FEV = \beta_0 + \beta_I \cdot I + \beta_U \cdot U + \beta_A \cdot A + \beta_S \cdot S + \beta_G \cdot G$$

+random variation

If the **three** levels/sources of **random** variation are **not** taken into account :

- The **precision** of the  $\beta_I$  and  $\beta_U$  are **highly overestimated**
- The **precision** of the  $\beta_A$  and  $\beta_S$  are **overestimated**
- The **estimates** of the  $\beta_I$  and  $\beta_U$  will be **biased** if the not all families are represented by the **same number of persons** and each person is measured the **same number of times**.
- The **estimates** of the  $\beta_A$  and  $\beta_S$  will be **biased** if the not all persons are measured the **same number of times**.

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Clustered data / data with several random components

$$FEV = \beta_0 + \beta_I \cdot I + \beta_U \cdot U + \beta_A \cdot A + \beta_S \cdot S + \beta_G \cdot G$$

$+F_f + P_{fp} + E_{fpd}$

$F_f$

$P_{fp}$

$E_{fpd}$

: Random family contribution

: Random person contribution

: Random day contribution

$\sigma_F^2$

$\sigma_P^2$

$\sigma_E^2$

$$\text{var}(FEV_{fpd}) = \sigma_F^2 + \sigma_P^2 + \sigma_E^2$$

Variance components

Assumed to be normal distributed

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Clustered data / data with several random components

Systematic part

$$FEV = \beta_0 + \beta_I \cdot I + \beta_U \cdot U + \beta_A \cdot A + \beta_S \cdot S + \beta_G \cdot G$$

Random part

$+F_f + P_{fp} + E_{fpd}$

$\beta_0, \beta_I, \beta_U, \beta_A, \beta_S$  and  $\beta_G$       Quantify the **systematic** variation

$\sigma_F^2, \sigma_P^2$  and  $\sigma_E^2$       Quantify the **random** variation

This is a:

- Variance component model
- Mixed model (both systematic and random variation)
- Multilevel model

The theory behind and the understanding of such models is well **established!!!**

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Clustered data / data with several random components

Dichotomous outcome

A different outcome:

$$H_{fpd} = \begin{cases} 1 & \text{if the person has hayfever} \\ 0 & \text{else} \end{cases}$$

A statistical model:

Systematic part

$$\text{logit}(H_{fpd} = 1) = \beta_0 + \beta_I \cdot I + \beta_U \cdot U + \beta_A \cdot A + \beta_S \cdot S + \beta_G \cdot G$$

Random part

$+F_f + P_{fp} + E_{fpd}$

This is not needed due to the binomial error

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Clustered data / data with several random components

Dichotomous outcome

$$\text{logit}(H_{fpd} = 1) = \beta_0 + \beta_I \cdot I + \beta_U \cdot U + \beta_A \cdot A + \beta_S \cdot S + \beta_G \cdot G$$

$+F_f + P_{fp}$

That is, an ordinary logistic regression + **random components**.

- A generalized linear mixed model
- A multilevel model for dichotomous outcome

Comments 1:

- It is **important** to include the **relevant random** components in the model.
- 'Multilevel models' is **essential** in medical/epidemiological research.

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Clustered data / data with several random components  
Dichotomous outcome

Comments 2:

- The theory and insight into the models for non-normal data are **not yet fully developed**.
- The main problem being that it is not known how to obtain **valid (unbiased) estimates**.
- Several software programs **falsely claim** to estimate the models. (SAS, STATA, SUDAAN, NLwin)
- The programs/algorithms are not able give 'the correct' estimates.

Advice:

At the moment, do not trust results based on multilevel models.

Wait and see, the statisticians might solve the problems.

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Clustered data / data with one random components  
Dichotomous outcome

If the models only involves **one random components**, e.g. **variation between families** or between GP's, then methods exists which can **adjust the standards errors**.

Remember that if the **data contains clusters**, then the precision of the estimates overestimated, that is the reported **standard errors is too small**.

So called **robust methods** or **sandwich estimates** of the standard errors will (try) adjust for this problem.

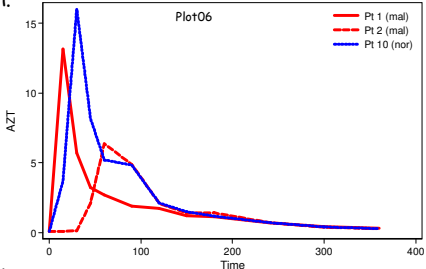
Only a **few** programs have this option - STATA does!

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Nonlinear regression models

Concentration in the blood of zidovudine (AZT) after administration of the drug.

One person with normal fat absorption and two with malabsorption.



Clearly non linear.

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Nonlinear regression models

The is **no way** that the above data can be described by a linear regression.

Furthermore **pharmaco kinetic** theory specify a simple model for the **expected concentration** as a function of time:

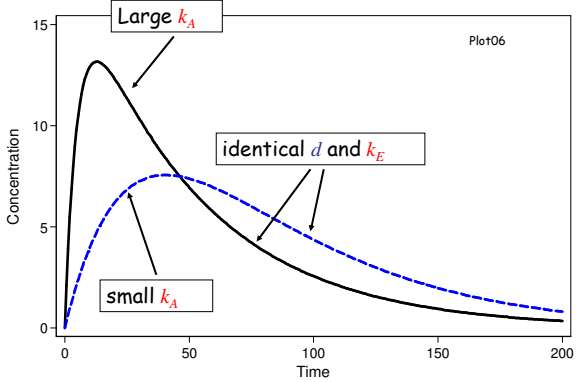
concentration(t) =  $d \frac{k_A}{k_A - k_E} (\exp(-k_E \cdot t) - \exp(-k_A \cdot t))$

Where:

- $d$  : dose (per kg bodyweight)
- $k_A$  : absorptionsrate
- $k_E$  : eliminationsrate

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Nonlinear regression models



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Nonlinear regression models

How do persons with malabsorption differ from normal in  $k_E$  and  $k_A$  ?

One type of analysis:

1. Fit this pharmaco-kinetic model to the data for each person.
2. Extract the estimates of  $k_E$  and  $k_A$  from each analysis.
3. Compare the distributions of  $k_E$  and  $k_A$  from the two types for persons.

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