

Multiple linear regression 2
Stata 11

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Categorical variables in regression models.
The changing reference level
Interaction/effect modification
Interaction between a categorical and continuous variable
Interaction between two categorical variables

Morten Frydenberg Linear and Logistic regression - Note 2.2 1

Categorical variable in regression models

The age distribution:

Let us divide *age* into three agegroups ,

0: $age \leq 40$, 1: $40 < age \leq 50$, 2: $50 < age$

and consider the new model

$$\ln(sbp) = \alpha_0 + \alpha_1 \cdot age1 + \alpha_2 \cdot age2 + \alpha_3 \cdot woman + \alpha_4 \cdot \ln\left(\frac{bmi}{25}\right) + E$$

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Categorical variable : age group 0 reference

$$\ln(sbp) = \alpha_0 + \alpha_1 \cdot age1 + \alpha_2 \cdot age2 + \alpha_3 \cdot woman + \alpha_4 \cdot \ln\left(\frac{bmi}{25}\right) + E$$

age1 is one if a person is in age group 1 and zero otherwise
age2 is one if a person is in age group 2 and zero otherwise

The expected $\ln(sbp)$ in the three age groups will be:

$age < 40: \ln(sbp) = \alpha_0 + \alpha_3 \cdot woman + \alpha_4 \cdot \ln(bmi/25)$

$40 \leq age < 50: \ln(sbp) = \alpha_0 + \alpha_1 + \alpha_3 \cdot woman + \alpha_4 \cdot \ln(bmi/25)$

$50 \leq age: \ln(sbp) = \alpha_0 + \alpha_2 + \alpha_3 \cdot woman + \alpha_4 \cdot \ln(bmi/25)$

We see that α_1 is the adjusted difference in $\ln(sbp)$ when comparing a person in the second group with one in the first group.

And α_2 is the adjusted difference in $\ln(sbp)$ when comparing a person in the third group with one in the first group.

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Categorical variable : age group 0 reference

$$\ln(sbp) = \alpha_0 + \alpha_1 \cdot age1 + \alpha_2 \cdot age2 + \alpha_3 \cdot woman + \alpha_4 \cdot \ln\left(\frac{bmi}{25}\right) + E$$

Finally we see that α_0 is the expected $\ln(sbp)$ for a man in the first (reference) age group, with $bmi=25$.

In most programs the model is fitted by first generating the grouping variable and then making the regression telling the program which variables are categorical.

In Stata this done is like this:

```
egen agegrp3=cut(age), at(0,40,50,120) 1tab1
regress lnsbp woman 1.agegrp3 lnBMI25
```

This is categorical

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Categorical variable : age group 0 reference

$$\ln(sbp) = \alpha_0 + \alpha_1 \cdot age1 + \alpha_2 \cdot age2 + \alpha_3 \cdot woman + \alpha_4 \cdot \ln\left(\frac{bmi}{25}\right) + E$$

OUTPUT:					
Source	SS	df	MS	Number of obs = 200	
Model	.980169926	4	.245042482	F(4, 195)	= 11.20
Residual	4.26524771	195	.021873065	Prob > F	= 0.0000
Total	5.24541764	199	.026358883	R-squared	= 0.1869
				Adj R-squared	= 0.1702
				Root MSE	= .1479

lnSBP	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
woman	.0035403	.0212026	0.17	0.868	-.0382757 .0453562
agegrp3					
1	.0715136	.0253373	2.82	0.005	.0215432 .121484
2	.130465	.0280521	4.65	0.000	.0751404 .1857895
lnBMI25	.2898622	.0772432	3.75	0.000	.1375229 .4422015
_cons	4.789641	.0224814	213.05	0.000	4.745303 4.833979

Note 0 is missing: it is base/reference group

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Categorical variable : age group 0 reference

$$\ln(sbp) = \alpha_0 + \alpha_1 \cdot age1 + \alpha_2 \cdot age2 + \alpha_3 \cdot woman + \alpha_4 \cdot \ln\left(\frac{bmi}{25}\right) + E$$

Adjusted difference between for a person in age group 1 compared to age group 0

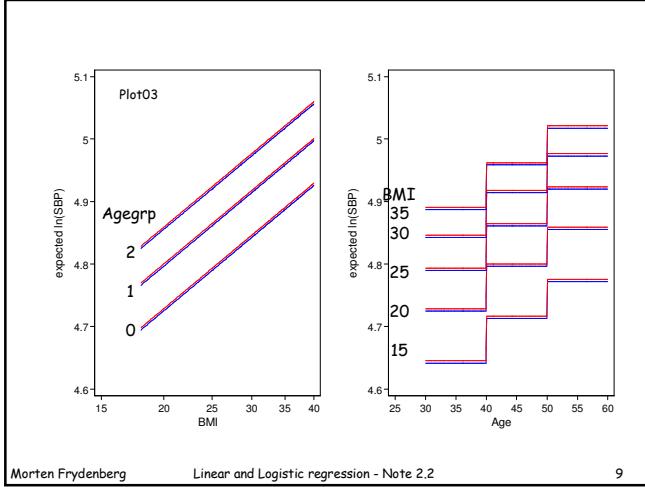
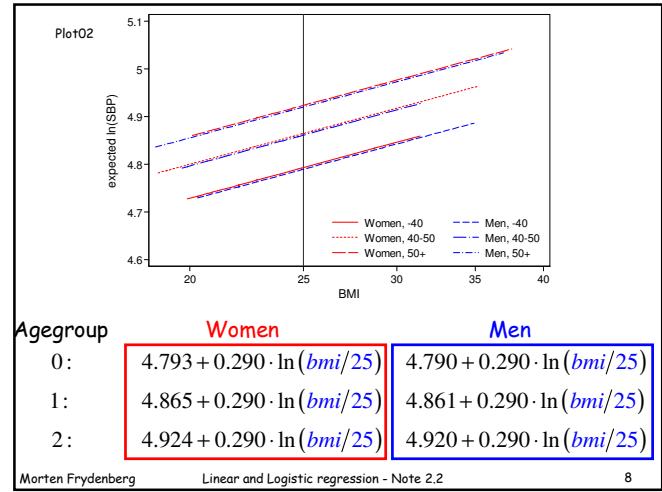
lnSBP	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
woman	.0035403	.0212026	0.17	0.868	-.0382757 .0453562
agegrp3					
1	.0715136	.0253373	2.82	0.005	.0215432 .121484
2	.130465	.0280521	4.65	0.000	.0751404 .1857895
lnBMI25	.2898622	.0772432	3.75	0.000	.1375229 .4422015
cons	4.789641	.0224814	213.05	0.000	4.745303 4.833979

Adjusted difference between for a person in age group 2 compared to age group 0

Expected value for a man in age group 0 with bmi=25.

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The expected values:		
$age < 40:$	$\ln(sbp) = \alpha_0 + \alpha_3 \cdot woman + \alpha_4 \cdot \ln(bmi/25)$	
$40 \leq age < 50:$	$\ln(sbp) = \alpha_0 + \alpha_1 + \alpha_3 \cdot woman + \alpha_4 \cdot \ln(bmi/25)$	
$50 \leq age:$	$\ln(sbp) = \alpha_0 + \alpha_2 + \alpha_3 \cdot woman + \alpha_4 \cdot \ln(bmi/25)$	
The estimates		
	lnSBP Coef.	
	woman .003540	3
1.agegrp3	.071514	1
2.agegrp3	.130465	2
lnBMI25	.289862	4
_cons	4.789641	0
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Categorical variable : age group 1 reference
 $\ln(sbp) = \gamma_0 + \gamma_1 \cdot age0 + \gamma_2 \cdot age2 + \gamma_3 \cdot woman + \gamma_4 \cdot \ln\left(\frac{bmi}{25}\right) + E$

$age0$ is one if a person is in age group 0 and zero otherwise
 $age2$ is one if a person is in age group 2 and zero otherwise

The expected $\ln(sbp)$ in the three age groups will be:

$age < 40:$ $\ln(sbp) = \gamma_0 + \gamma_1 + \gamma_3 \cdot woman + \gamma_4 \cdot \ln(bmi/25)$

$40 \leq age < 50:$ $\ln(sbp) = \gamma_0 + \gamma_3 \cdot woman + \gamma_4 \cdot \ln(bmi/25)$

$50 \leq age:$ $\ln(sbp) = \gamma_0 + \gamma_2 + \gamma_3 \cdot woman + \gamma_4 \cdot \ln(bmi/25)$

We see that γ_1 is the adjusted difference in $\ln(sbp)$ when comparing a person in the first group with one in the second group.

And γ_2 is the adjusted difference in $\ln(sbp)$ when comparing a person in the third group with one in the second group.

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Categorical variable : age group 1 reference		
$\ln(sbp) = \gamma_0 + \gamma_1 \cdot age0 + \gamma_2 \cdot age2 + \gamma_3 \cdot woman + \gamma_4 \cdot \ln\left(\frac{bmi}{25}\right) + E$		
Finally we see that γ_1 is the expected $\ln(sbp)$ for a man in the second (reference) age group, with $bmi=25$.		
Many programs (but regression in SPSS) let you choose the reference group		
In Stata 11 this is done like this:		
<code>regress lnSBP woman b1.agegrp3 lnBMI25</code>		
This is categorical with base/reference set to 1		

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Categorical variable : age group 1 reference		
$\ln(sbp) = \gamma_0 + \gamma_1 \cdot age0 + \gamma_2 \cdot age2 + \gamma_3 \cdot woman + \gamma_4 \cdot \ln\left(\frac{bmi}{25}\right) + E$		
Source SS df MS		Number of obs = 200
Model .980169926 4 .245042482		F(4, 195) = 11.20
Residual 4.26524771 195 .021873065		Prob > F = 0.0000
Total 5.24541764 199 .026358883		R-squared = 0.1869
		Adj R-squared = 0.1702
		Root MSE = .1479
lnSBP Coef. Std. Err. t P> t [95% Conf. Interval]		
woman .0035403 .0212026 0.17 0.868 -.0382757 .0453562		
agegrp3		
0 -.0715136 .0253373 -2.82 0.005 -.121484 -.0215432		
2 .0589513 .0263496 2.24 0.026 .0069846 .1109181		
lnBMI25 .2898622 .0772432 3.75 0.000 .1375229 .4422015		
_cons 4.861154 .0207406 234.38 0.000 4.82025 4.902059		
Note 1 is missing: it is base/reference group		

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Categorical variable : age group 1 reference

$$\ln(sbp) = \gamma_0 + \gamma_1 \cdot \text{age0} + \gamma_2 \cdot \text{age2} + \gamma_3 \cdot \text{woman} + \gamma_4 \cdot \ln\left(\frac{\text{bmi}}{25}\right) + E$$

Adjusted difference between for a person in age group 0 compared to age group 1

lnSBP	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
woman	.0035403	.0212026	0.17	0.868	-.0382757 .0453562	
agegrp3	0	-.0715136	.0253373	-2.82	0.005	-.121484 -.0215432
2	.0589513	.0263496	2.24	0.026	.0069846 .1109181	
lnBMI25	.2898622	.0772432	3.75	0.000	.1375229 .4422015	
_cons	4.861154	.0207406	234.38	0.000	4.82025 4.902059	

Adjusted difference between for a person in age group 2 compared to age group 1

Expected value for a man in age group 1 with bmi=25.

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Categorical variable: Comparing two parameterizations

$$\ln(sbp) = \alpha_0 + \alpha_1 \cdot \text{age1} + \alpha_2 \cdot \text{age2} + \alpha_3 \cdot \text{woman} + \alpha_4 \cdot \ln\left(\frac{\text{bmi}}{25}\right) + E$$

$$\ln(sbp) = \gamma_0 + \gamma_1 \cdot \text{age0} + \gamma_2 \cdot \text{age2} + \gamma_3 \cdot \text{woman} + \gamma_4 \cdot \ln\left(\frac{\text{bmi}}{25}\right) + E$$

age group 0: $\alpha_0 = \gamma_0 + \gamma_1$
age group 1: $\alpha_0 + \alpha_1 = \gamma_0$
age group 2: $\alpha_0 + \alpha_2 = \gamma_0 + \gamma_2$

$$\alpha_0 = \gamma_0 + \gamma_1 \quad \gamma_0 = \alpha_0 + \alpha_1$$

$$\alpha_1 = -\gamma_1 \quad \gamma_1 = -\alpha_1$$

$$\alpha_2 = \gamma_2 - \gamma_1 \quad \gamma_2 = \alpha_2 - \alpha_1$$

$$\alpha_3 = \gamma_3 \quad \gamma_3 = \alpha_3$$

$$\alpha_4 = \gamma_4 \quad \gamma_4 = \alpha_4$$

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Categorical variable: Comparing two parameterizations

The estimates:
age group 0 reference **age group 1 reference**

Root MSE	= .1479	Root MSE	= .1479
lnSBP	coef. [95% CI]	lnSBP	coef. [95% CI]
woman	.0035 -.0382 .0453	woman	.0035 -.0382 .0453
agegrp3	0 .0715 .0215 .1214 2 .1304 .0751 .1857 lnBMI25 .2898 .1375 .4422 _cons 4.7896 4.745 4.833	agegrp3	0 -.0715 -.1214 -.0215 2 .0589 .0069 .1109 lnBMI25 .2898 .1375 .4422 _cons 4.8611 4.820 4.902

Note, the estimates fulfil the same equations.
The interpretation of the "agegrp3 2 line" and "cons line" are altered!!!!!!!
Always remember: what is the reference group!

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Categorical variable : age group 1 reference

lnSBP	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
woman	.0035403	.0212026	0.17	0.868	-.0382757 .0453562	
agegrp3	0	-.0715136	.0253373	-2.82	0.005	-.121484 -.0215432
2	.0589513	.0263496	2.24	0.026	.0069846 .1109181	
lnBMI25	.2898622	.0772432	3.75	0.000	.1375229 .4422015	
_cons	4.861154	.0207406	234.38	0.000	4.82025 4.902059	

Two test:
One testing no difference between age group 0 and 1.
One testing no difference between age group 2 and 1.
Can we get one test testing no difference between age groups?
An F-test in Stata: `testparm i.agegrp`

$(1) \text{agegrp3} = 0$
 $(2) \text{agegrp3} = 0$
 $F(2, 195) = 10.93$
 $\text{Prob} > F = 0.0000$

Highly significant

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Linear and Logistic regression - Note 2.2

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Interactions/effectmodification

$$\ln(sbp) = \gamma_0 + \gamma_1 \cdot \text{age0} + \gamma_2 \cdot \text{age2} + \gamma_3 \cdot \text{woman} + \gamma_4 \cdot \ln\left(\frac{\text{bmi}}{25}\right) + E$$

One of the central assumptions was "no effect modification".
E.g. in the model above the "effect" of age, sex and bmi did not depend on the value of each other.
One can introduce effect modification between a categorical variable and another variable.
Here we first will look at agegrp3 and lnBMI25.
The effect modification will be that the coefficient to lnBMI25 depend on age group.
That is, we will allow different effect of bmi in the different age groups.

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Linear and Logistic regression - Note 2.2

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Interactions/effectmodification

$$\ln(sbp) = \omega_0 + \omega_1 \cdot \text{age0} + \omega_2 \cdot \text{age2} + \omega_3 \cdot \text{woman}$$

$$+ \omega_4 \cdot \ln\left(\frac{\text{bmi}}{25}\right) + \omega_5 \cdot \text{age0} \cdot \ln\left(\frac{\text{bmi}}{25}\right) + \omega_6 \cdot \text{age2} \cdot \ln\left(\frac{\text{bmi}}{25}\right) + E$$

$\text{age} \leq 40:$ $\ln(sbp) = (\omega_0 + \omega_1) + (\omega_4 + \omega_5) \cdot \ln\left(\frac{\text{bmi}}{25}\right) + \omega_3 \cdot \text{woman}$

$40 < \text{age} \leq 50:$ $\ln(sbp) = \omega_0 + \omega_4 \cdot \ln\left(\frac{\text{bmi}}{25}\right) + \omega_3 \cdot \text{woman}$

$50 < \text{age}:$ $\ln(sbp) = (\omega_0 + \omega_1) + (\omega_4 + \omega_6) \cdot \ln\left(\frac{\text{bmi}}{25}\right) + \omega_3 \cdot \text{woman}$

ω_1 is the difference between the constant for age group 0 and reference group.
 ω_5 is the difference between the coefficient to lnBMI25 for age group 0 and reference group.

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Interactions/effect modification						
regress lnSBP woman b1.agegrp3#c.lnBMI25						
Source	SS	df	MS	Number of obs = 200		
Model	.994860827	6	.165810138	F(6, 193) = 7.53		
Residual	4.25055681	193	.02202361	Prob > F = 0.0000		
Total	5.24541764	199	.026358883	R-squared = 0.1897		
				Adj R-squared = 0.1645		
				Root MSE = .1484		
lnSBP	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
woman	.0076438	.0219199	0.35	0.728	-.0355896	.0508772
agegrp3						
0	-.0708045	.0261198	-2.71	0.007	-.1223213	-.0192877
2	.0631082	.0270342	2.33	0.021	.0097877	.1164287
lnBMI25	.3155479	.1222905	2.58	0.011	.0743505	.5567453
agegrp3#c.lnBMI25						
0	.0429736	.1912373	0.22	0.822	-.3342099	.420157
2	-.1165375	.1855477	-0.63	0.531	-.4824991	.2494242
_cons	4.859743	.0214122	226.96	0.000	4.817511	4.901975

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Linear and Logistic regression - Note 2.2

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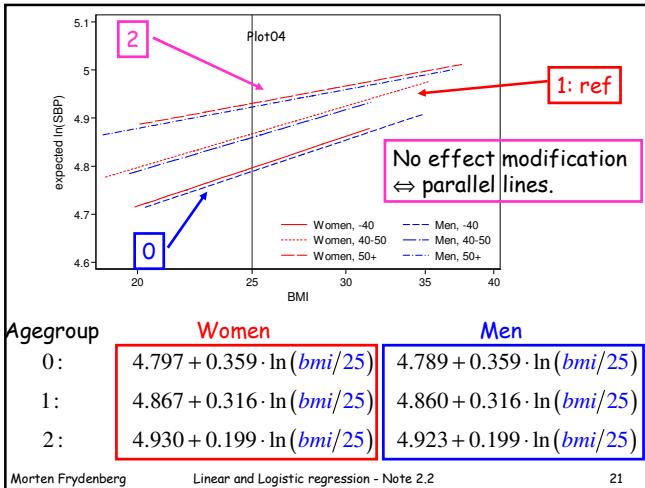
Interactions/effect modification						
Ref: constant and 'slope' in reference group						
O: difference in constant and slope compared to reference						
2: difference in constant and slope compared to reference						
lnSBP	Coef.	Std. E.	t	P> t	[95% Conf. Interval]	
woman	.0076438	.0219199	0.35	0.728	-.0355896	.0508772
agegrp3						
0	-.0708045	.0261198	-2.71	0.007	-.1223213	-.0192877
2	.0631082	.0270342	2.33	0.021	.0097877	.1164287
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agegrp3#c.lnBMI25						
0	.0429736	.1912373	0.22	0.822	-.3342099	.420157
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_cons	4.859743	.0214122	226.96	0.000	4.817511	4.901975

Note the larger standard errors

Based on the estimates one can find the six "dose-response" curves:

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Linear and Logistic regression - Note 2.2

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Interactions/effect modification						
lnSBP	Coef.	Std. E.	t	P> t	[95% Conf. Interval]	
woman	.0076438	.0219199	0.35	0.728	-.0355896	.0508772
agegrp3						
0	-.0708045	.0261198	-2.71	0.007	-.1223213	-.0192877
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_cons	4.859743	.0214122	226.96	0.000	4.817511	4.901975

Two tests:

One testing differences between "slope" in age group 0 and 1.

One testing differences between "slope" in age group 2 and 1.

One test testing no difference between age groups!

A F-test in Stata: `testparm i.agegrp3#c.lnBMI25`

```
( 1) 0.agegrp3#c.lnBMI25 = 0
( 2) 2.agegrp3#c.lnBMI25 = 0
F( 2, 193) = 0.33
Prob > F = 0.7168
```

Non-significant

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Interactions/effect modification						
The test of no interaction was non-significant.						
But look at the confidence interval for the difference in slope for between age group 2 and group 1!						
lnSBP	Coef.	Std. E.	t	P> t	[95% Conf. Interval]	
woman	.0076438	.0219199	0.35	0.728	-.0355896	.0508772
agegrp3						
0	-.0708045	.0261198	-2.71	0.007	-.1223213	-.0192877
2	.0631082	.0270342	2.33	0.021	.0097877	.1164287
lnBMI25	.3155479	.1222905	2.58	0.011	.0743505	.5567453
agegrp3#c.lnBMI25						
0	.0429736	.1912373	0.22	0.822	-.3342099	.420157
2	-.1165375	.1855477	-0.63	0.531	-.4824991	.2494242
_cons	4.859743	.0214122	226.96	0.000	4.817511	4.901975

It is very wide!!! We know very little about this difference!

The test for no interaction has very low power!!!

The data have very little information on whether there is effect modification.

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Linear and Logistic regression - Note 2.2

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Interaction between age group and sex						
regress lnSBP 1nBMI25 b1.agegrp3##b1.sex Male sex=-1						
lnSBP	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
1nBMI25	.2265018	.0774898	2.92	0.004	.0736662	.3793374
agegrp3						
0	-.0426734	.0377096	-1.13	0.259	-.1170493	.0317025
2	-.0025412	.0365457	-0.07	0.945	-.0746215	.0695391
2.sex	-.0210869	.0322283	-0.65	0.514	-.0846518	.042478
agegrp3#sex						
0 2	-.0548967	.0501668	-1.09	0.275	-.1538422	.0440488
2 2	.133379	.0501308	2.66	0.008	.0345043	.2322536
_cons	4.873442	.0251767	193.57	0.000	4.823786	4.923099

. testparm i.agegrp3#sex

```
( 1) 0.agegrp3#sex = 0
( 2) 2.agegrp3#sex = 0
```

```
F( 2, 193) = 6.26
Prob > F = 0.0023
```

Highly significant

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Interaction between age group and sex						
Differences between age groups among men are small						
<hr/>						
lnSBP	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
lnBMI25	.2265018	.0774898	2.92	0.004	.0736662	.3793374
agegrp3						
0	-.0426734	.0377096	-1.13	0.259	-.1170493	.0317025
2	-.0025412	.0365457	-0.07	0.945	-.0746215	.0695391
2.sex	-.0210869	.0322283	-0.65	0.514	-.0846518	.042478
agegrp3#sex						
0 2	-.0548967	.0501668	-1.09	0.275	-.1538422	.0440488
2 2	.133379	.0501308	2.66	0.008	.0345043	.2322536
_cons	4.873442	.0251767	193.57	0.000	4.823786	4.923099
<hr/>						
Women age group 1: 4.873–0.021=4.852						
Women age group 0: 4.873–0.021=0.042–0.055=4.755						
Women age group 2: 4.873–0.021=0.003+0.133=4.982						
<hr/>						
Large differences in the age groups among women.						

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Linear and Logistic regression - Note 2.2

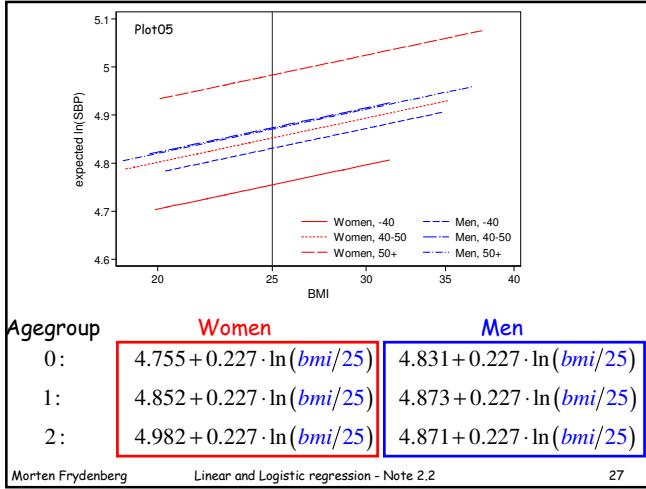
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Interaction between age group and sex						
Using women as reference: b2.sex						
Large differences between age groups among women						
<hr/>						
lnSBP	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
lnBMI25	.2265018	.0774898	2.92	0.004	.0736662	.3793374
agegrp3						
0	-.0975701	.0328469	-2.97	0.003	-.162355	-.0327852
2	.1308378	.0354804	3.69	0.000	.0608587	.2008168
1.sex	.0210869	.0322283	0.65	0.514	-.042478	.0846518
agegrp3#sex						
0 1	.0548967	.0501668	1.09	0.275	-.0440488	.1538422
2 1	-.133379	.0501308	-2.66	0.008	-.2322536	-.0345043
_cons	4.852355	.0205502	236.12	0.000	4.811824	4.892887
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Linear and Logistic regression - Note 2.2

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