

Extensions to linear and logistic regressions
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Conditional logistic regression

- When?
- What?
- How?

Other methods for analyzing binary data

- Models for **relative risks**
- Models for **risk differences**

Clustered data / data with several random components

- Continuous outcome**
- Dichotomous outcome**

Clustered binary data with one random components

Nonlinear regression models

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Conditional logistic regression
When

Used in two situations:

1. **Matched studies** (binary response).
2. **Unmatched studies with a confounder with many distinct values.**

In 1. the models correspond to **the way data was collected**.

In 2. the method adjust for a '**mathematical**' flaw in the unconditional method.

An example of situation 2. the confounder is " **kommune**" having 275 distinct values.

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Conditional logistic regression
What

The logistic regression model (outcome disease yes/no):

$$\ln(\text{odds}) = \alpha + \sum_{i=1}^k (\beta_i \cdot x_i)$$

Suppose the model above hold in each strata:

$$\ln(\text{odds}) = \alpha_s + \sum_{i=1}^k (\beta_i \cdot x_i)$$

ln(odds) in reference ln(odds ratios)
different in each strata the same in each strata

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Conditional logistic regression
What

$$\ln(\text{odds}) = \alpha_s + \sum_{i=1}^k (\beta_i \cdot x_i)$$

ln(odds) different in each strata
We are not interested in these !

In a **matched study** these are '**controlled**'.

In a **conditional logistic regression** one '**condition on the odds in each strata**', i.e. these case/control ratio.

In the conditional model the **α 's disappear** !

The **β 's**, the log OR's, are still in and **can be estimated**.

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Conditional logistic regression How

It is easy !

You need a statistical software package.

A package made for **research in epidemiology**

Not in social science

Not **SPSS**

But **STATA, EPICURE, EPILOG, EGRET, EPIINFO(2000) and SAS** can do it.

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Conditional logistic regression How

An example using **STATA**

A study of cancer in the oral cavity

Matched on **gender** and **10 years age groups**

Ten strata (**genage**)

Here we focus on

textile-worker and

life time consumption of alcohol (three groups)

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Conditional logistic regression How

logistic regression in **STATA**

xi:clogit cancer textile i.alkcon i.genage

Part of the output:

| | cancer | Coef. | Std. Err. | z | P> z | CI |
|--------------------|--------|---------|-----------|--------|-------|----------------|
| textile | | .5022 | .4141 | 1.213 | 0.225 | -.3094 1.3139 |
| <u>i.alkcon_1</u> | | .4628 | .2823 | 1.639 | 0.101 | -.0905 1.0163 |
| <u>i.alkcon_2</u> | | 2.7165 | .3232 | 8.404 | 0.000 | 2.0829 3.3501 |
| <u>i.genage_2</u> | | .2450 | 1.2514 | 0.196 | 0.845 | -2.2075 2.6977 |
| <u>i.genage_3</u> | | -.4940 | .5503 | -0.898 | 0.369 | -1.5726 .5846 |
| <u>i.genage_4</u> | | 1798 | .6406 | 0.281 | 0.779 | -1.0758 1.4353 |
| <u>i.genage_5</u> | | -.2899 | .5482 | -0.529 | 0.597 | -1.3644 .7844 |
| <u>i.genage_6</u> | | .2127 | .6262 | 0.340 | 0.734 | -1.0147 1.4401 |
| <u>i.genage_7</u> | | -.2305 | .5355 | -0.431 | 0.667 | -1.2802 .8190 |
| <u>i.genage_8</u> | | .5507 | .5263 | 1.046 | 0.295 | -4.4809 1.5825 |
| <u>i.genage_9</u> | | .0315 | .5884 | 0.054 | 0.957 | -1.1217 1.1847 |
| <u>i.genage_10</u> | | .5572 | .5595 | 0.996 | 0.319 | -.53954 1.6539 |
| const | | -1.4692 | .4762 | -3.085 | 0.002 | -2.4027 -.5356 |

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Conditional logistic regression in **STATA**

The syntax:

xi:clogit cancer textile i.alkcon, group(genage)

Part of the output:

| | cancer | Coef. | Std. Err. | z | P> z | CI |
|-------------------|--------|-------|-----------|-------|-------|---------------|
| textile | | .4929 | .4103 | 1.201 | 0.230 | -.3112 1.2971 |
| <u>i.alkcon_1</u> | | .452 | .27923 | 1.621 | 0.105 | -.094 .9999 |
| <u>i.alkcon_2</u> | | 2.660 | .31936 | 8.332 | 0.000 | 2.034 3.2868 |

xi:clogit cancer textile i.alkcon, group(genage) or

| | cases | Odds Ratio | Std. Err. | z | P> z | [95% Conf. Interval] |
|-------------------|-------|------------|-----------|------|-------|----------------------|
| textile | | 1.63708 | .6717022 | 1.20 | 0.230 | .732517 3.658661 |
| <u>i.alkcon_1</u> | | 1.572508 | .4390957 | 1.62 | 0.105 | .909724 2.718168 |
| <u>i.alkcon_2</u> | | 14.30908 | 4.569879 | 8.33 | 0.000 | 7.651811 26.75835 |

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Other methods to analysis of binary response data

Relative Risk models

Logistic regression model focus on the Odds Ratios

This is the correct thing to do in **case-control** studies.

In **follow-up studies** **Relative Risk** is often the appropriate measure of association, (personal risk).

I.e. a model like this might be more relevant:

$$\Pr(\text{event}) = p_0 \times RR_1 \times RR_2 \times RR_3$$

$$\ln\{\Pr(\text{event})\} = \ln(p_0) + \ln(RR_1) + \ln(RR_2) + \ln(RR_3)$$

$$\ln\{\Pr(\text{event given the covariates})\} = \alpha + \sum_{i=1}^p (\beta_i \cdot x_i)$$

That is linear on **log-probability scale**

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Other methods to analysis of binary response data

Relative Risk models

$$\ln\{\Pr(\text{event given the covariates})\} = \alpha + \sum_{i=1}^p (\beta_i \cdot x_i)$$

Such a model **modelling the relative risk** can easily be fitted by many programs (not SPSS).

Logistic regression in STATA:

xi: logit obese age i.sex

or

xi: glm obese age i.sex, fam(bin) link(logit)

Relative risk model:

xi: glm obese age i.sex, fam(bin) link(id)

The **link** is **log** instead of **logit**

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Other methods to analysis of binary response data

Risk difference models

Logistic regression model focus on the Odds Ratios

This is the correct thing to do in **case-control** studies.

In **follow-up studies** **Risk Difference** is often the appropriate measure of association, (community effect).

I.e. a model like this might be more relevant:

$$\Pr(\text{event}) = p_0 + RD_1 + RD_2 + RD_3$$

$$\Pr(\text{event given the covariates}) = \alpha + \sum_{i=1}^p (\beta_i \cdot x_i)$$

That is linear on **probability scale**

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Other methods to analysis of binary response data

Risk difference models

$$\Pr(\text{event given the covariates}) = \alpha + \sum_{i=1}^p (\beta_i \cdot x_i)$$

Such a model **modelling the risk difference** can easily be fitted by many programs (not SPSS).

Logistic regression in STATA:

xi: logit obese age i.sex

or

xi: glm obese age i.sex, fam(bin) link(logit)

Risk difference model:

xi: glm obese age i.sex, fam(bin) link(id)

The **link** is **identity** instead of **logit**

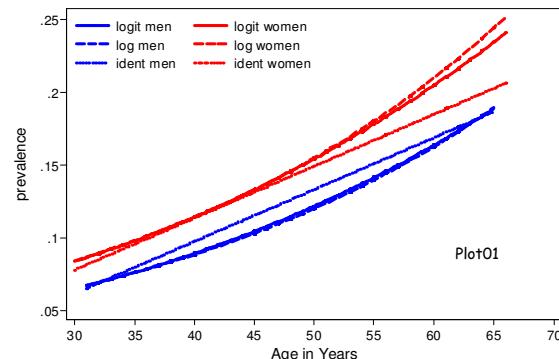
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Other methods to analysis of binary response data

Three different links for *Obese* "=" *sex* "+" *age*



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Other methods to analysis of binary response data Problems

$$\Pr(\text{event}) = p_0 \times RR_1 \times RR_2 \times RR_3$$

As the relative risk can be **larger** than one
the product might be **larger** than one !

$$\Pr(\text{event}) = p_0 + RD_1 + RD_2 + RD_3$$

The sum might **negative** and be **larger than one** !

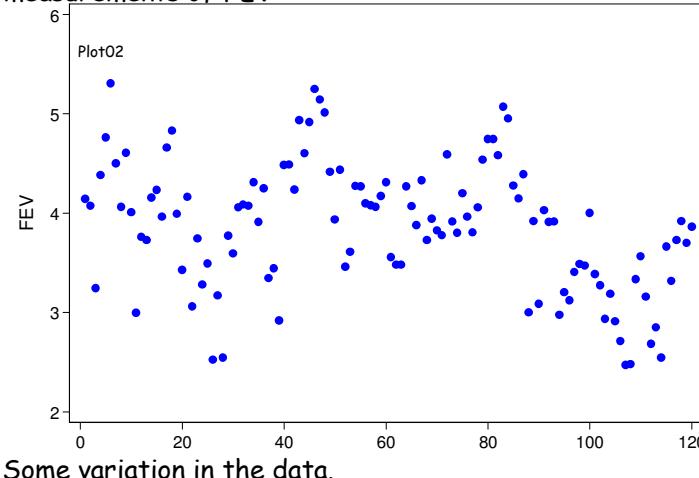
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Clustered data / data with several random components

120 measurements of FEV:



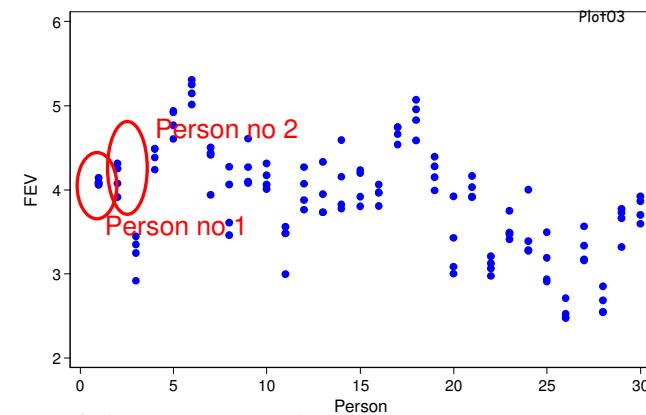
Some variation in the data.

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Clustered data / data with several random components But it is on only 30 persons:



Some of the variation is due to **variation between persons** and some within person.

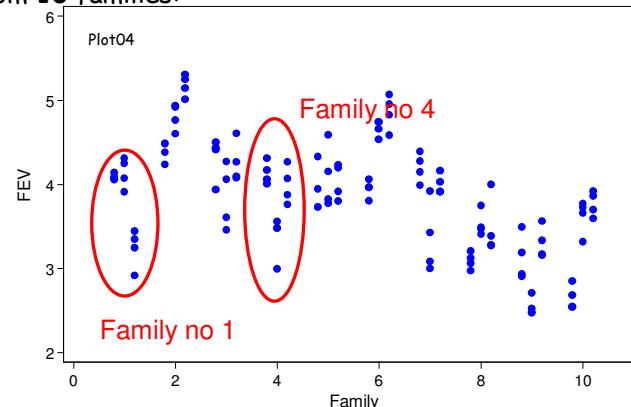
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Clustered data / data with several random components

From 10 families:



Some of the variation between persons is due to variation between families and some within family.

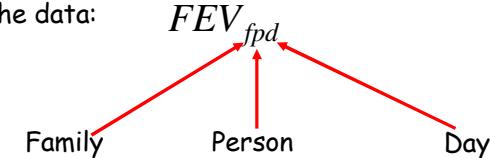
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Clustered data / data with several random components

Structure of the data:



Three sources of random variation:

Variation between families

Variation between persons (variation within family)

Variation between days (variation within person)

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Clustered data / data with several random components

Factors of interest:

| | | |
|--------------|--------|--|
| household | Income | Constant within family |
| Urbanization | | Constant within family |
| Age | | Constant within person; varies within family |
| Sex | | Constant within person; varies within family |
| Grass pollen | | Constant within day; varies within person |

A model:

$$FEV = \beta_0 + \beta_I \cdot I + \beta_U \cdot U + \beta_A \cdot A + \beta_S \cdot S + \beta_G \cdot G + \text{random variation}$$

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Clustered data / data with several random components

$$FEV = \beta_0 + \beta_I \cdot I + \beta_U \cdot U + \beta_A \cdot A + \beta_S \cdot S + \beta_G \cdot G + \text{random variation}$$

If the three levels/sources of random variation are not taken into account :

- The precision of the β_I and β_U are highly overestimated
- The precision of the β_A and β_S are overestimated
- The estimates of the β_I and β_U will be biased if the not all families are represented by the same number of persons and each person is measured the same number of times.
- The estimates of the β_A and β_S will be biased if the not all persons are measured the same number of times.

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Clustered data / data with several random components

$$FEV = \beta_0 + \beta_I \cdot I + \beta_U \cdot U + \beta_A \cdot A + \beta_S \cdot S + \beta_G \cdot G + F_f + P_{fp} + E_{fpd}$$

variance

| | | |
|-----------|------------------------------|--------------|
| F_f | : Random family contribution | σ_F^2 |
| P_{fp} | : Random person contribution | σ_P^2 |
| E_{fpd} | : Random day contribution | σ_E^2 |

$$\text{var}(FEV_{fpd}) = \sigma_F^2 + \sigma_P^2 + \sigma_E^2$$

Variance components

Assumed to be normal distributed

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Clustered data / data with several random components

Systematic part

$$FEV = \beta_0 + \beta_I \cdot I + \beta_U \cdot U + \beta_A \cdot A + \beta_S \cdot S + \beta_G \cdot G + F_f + P_{fp}$$

Random part

$\beta_0, \beta_I, \beta_U, \beta_A, \beta_S$ and β_G Quantify the **systematic** variation

σ_F^2, σ_P^2 and σ_E^2 Quantify the **random** variation

This is a:

- **Variance component model**
- **Mixed model** (both systematic and random variation)
- **Multilevel model**

The theory behind and the understanding of such models is well **established!!!**

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Clustered data / data with several random components

Dichotomous outcome

A different outcome:

$$H_{fpd} = \begin{cases} 1 & \text{if the person has hayfewer} \\ 0 & \text{else} \end{cases}$$

A statistical model:

Systematic part

$$\text{logit}(H_{fpd} = 1) = \beta_0 + \beta_I \cdot I + \beta_U \cdot U + \beta_A \cdot A + \beta_S \cdot S + \beta_G \cdot G + F_f + P_{fp}$$

Random part

This is not needed due to the binomial error

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Clustered data / data with several random components

Dichotomous outcome

$$\text{logit}(H_{fpd} = 1) = \beta_0 + \beta_I \cdot I + \beta_U \cdot U + \beta_A \cdot A + \beta_S \cdot S + \beta_G \cdot G + F_f + P_{fp}$$

That is, an ordinary logistic regression + **random components**.

- **A generalized linear mixed model**
- **A multilevel model for dichotomous outcome**

Comments 1:

- It is **important** to include the **relevant random components** in the model.
- 'Multilevel models' is **essential** in medical/epidemiological research.

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Clustered data / data with several random components Dichotomous outcome

Comments 2:

- The theory and insight into the models for non-normal data are **not yet fully developed**.
- The main problem being that it is not known how to obtain **valid (unbiased) estimates**.
- Several software programs **falsey claim** to estimate the models. (SAS, STATA, SUDAAN, NLwin)
- The programs/algorithms are not able give 'the correct' estimates.

Advice:

At the moment, do not trust results based on multilevel models.

Wait and see, the statisticians **might** solve the problems.

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Clustered data / data with one random components Dichotomous outcome

If the models only involves **one random components**, e.g. **variation between families** or between GP's,

then methods exists which can **adjust the standards errors**.

Remember that if the **data contains clusters**, then the precision of the estimates overestimated, that is the reported **standard errors is too small**.

So called **robust methods** or **sandwich estimates** of the standard errors will (try) adjust for this problem.

Only a **few** programs have this option - STATA does!

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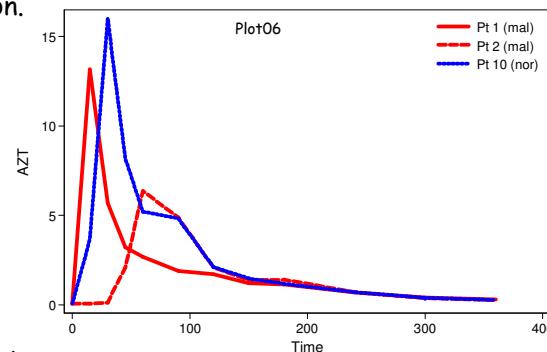
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Nonlinear regression models

Concentration in the blood of zidovudine (AZT) after administration of the drug.

One person with normal fat absorption and two with malabsorption.



Clearly non linear.

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Nonlinear regression models

The is **no way** that the above data can be described by a linear regression.

Furthermore **pharmaco kinetic** theory specify a simple model for the **expected concentration** as a function of time:

$$\text{concentration}(t) = d \frac{k_A}{k_A - k_E} (\exp(-k_E \cdot t) - \exp(-k_A \cdot t))$$

Where:

d : dose (per kg bodyweight)

k_A : absorptionsrate

k_E : eliminationsrate

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