

Extensions to linear and logistic regressions
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Conditional logistic regression

- When?
- What?
- How?

Other methods for analyzing binary data

- Models for **relative risks**
- Models for **risk differences**

Clustered data / data with several random components

- Continuous outcome**
- Dichotomous outcome**

Clustered binary data with one random components

Nonlinear regression models

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Conditional logistic regression
When

Used in two situations:

1. **Matched studies** (binary response).
2. **Unmatched studies with a confounder with many distinct values.**

In 1. the models correspond to **the way data was collected**.
 In 2. the method adjust for a '**mathematical**' flaw in the unconditional method.

An example of situation 2. the confounder is " **kommune**" having 275 distinct values.

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Conditional logistic regression
What

The logistic regression model (outcome disease yes/no):

$$\ln(\text{odds}) = \alpha + \sum_{i=1}^k (\beta_i \cdot x_i)$$

Suppose the model above hold in each strata:

$$\ln(\text{odds}) = \alpha_s + \sum_{i=1}^k (\beta_i \cdot x_i)$$

ln(odds) in reference ln(odds ratios)
different in each strata the same in each strata

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Conditional logistic regression
What

$$\ln(\text{odds}) = \alpha_s + \sum_{i=1}^k (\beta_i \cdot x_i)$$

ln(odds) different in each strata
We are not interested in these !

In a **matched study** these are '**controlled**'.
 In a **conditional logistic regression** one '**condition on the odds in each strata**', i.e. these case/control ratio.
 In the **conditional model** the **α 's disappear**!
 The **β 's**, the log OR's, are still in and **can be estimated**.

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Conditional logistic regression How

It is easy !

You need a statistical software package.

A package made for **research in epidemiology**

Not in social science

Not **SPSS**

But **STATA, EPICURE, EPILOG, EGRET, EPIINFO(2000) and SAS** can do it.

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5

Conditional logistic regression How

An example using **STATA**

A study of cancer in the oral cavity

Matched on **gender** and **10 years age groups**

Ten strata (**genage**)

Here we focus on

textile-worker and

life time consumption of alcohol (three groups)

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6

Conditional logistic regression How

logistic regression in **STATA**

xi:clogit cancer textile i.alkcon i.genage

Part of the output:

	cancer	Coef.	Std. Err.	z	P> z	CI
textile		.5022	.4141	1.213	0.225	-.3094 1.3139
<u>i.alkcon_1</u>		.4628	.2823	1.639	0.101	-.0905 1.0163
<u>i.alkcon_2</u>		2.7165	.3232	8.404	0.000	2.0829 3.3501
<u>i.genage_2</u>		.2450	1.2514	0.196	0.845	-2.2075 2.6977
<u>i.genage_3</u>		-.4940	.5503	-0.898	0.369	-1.5726 .5846
<u>i.genage_4</u>		1798	.6406	0.281	0.779	-1.0758 1.4353
<u>i.genage_5</u>		-.2899	.5482	-0.529	0.597	-1.3644 .7844
<u>i.genage_6</u>		.2127	.6262	0.340	0.734	-1.0147 1.4401
<u>i.genage_7</u>		-.2305	.5355	-0.431	0.667	-1.2802 .8190
<u>i.genage_8</u>		.5507	.5263	1.046	0.295	-4.4809 1.5825
<u>i.genage_9</u>		.0315	.5884	0.054	0.957	-1.1217 1.1847
<u>i.genage_10</u>		.5572	.5595	0.996	0.319	-.53954 1.6539
const		-1.4692	.4762	-3.085	0.002	-2.4027 -.5356

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7

Conditional logistic regression in **STATA**

The syntax:

xi:clogit cancer textile i.alkcon, group(genage)

Part of the output:

	cancer	Coef.	Std. Err.	z	P> z	CI
textile		.4929	.4103	1.201	0.230	-.3112 1.2971
<u>i.alkcon_1</u>		.452	.27923	1.621	0.105	-.094 .9999
<u>i.alkcon_2</u>		2.660	.31936	8.332	0.000	2.034 3.2868

xi:clogit cancer textile i.alkcon, group(genage) or

	cases	Odds Ratio	Std. Err.	z	P> z	[95% Conf. Interval]
textile		1.63708	.6717022	1.20	0.230	.732517 3.658661
<u>i.alkcon_1</u>		1.572508	.4390957	1.62	0.105	.909724 2.718168
<u>i.alkcon_2</u>		14.30908	4.569879	8.33	0.000	7.651811 26.75835

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8

Other methods to analysis of binary response data

Relative Risk models

Logistic regression model focus on the Odds Ratios

This is the correct thing to do in **case-control** studies.

In **follow-up studies** **Relative Risk** is often the appropriate measure of association, (personal risk).

I.e. a model like this might be more relevant:

$$\Pr(\text{event}) = p_0 \times RR_1 \times RR_2 \times RR_3$$

$$\ln\{\Pr(\text{event})\} = \ln(p_0) + \ln(RR_1) + \ln(RR_2) + \ln(RR_3)$$

$$\ln\{\Pr(\text{event given the covariates})\} = \alpha + \sum_{i=1}^p (\beta_i \cdot x_i)$$

That is linear on **log-probability scale**

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9

Other methods to analysis of binary response data

Relative Risk models

$$\ln\{\Pr(\text{event given the covariates})\} = \alpha + \sum_{i=1}^p (\beta_i \cdot x_i)$$

Such a model **modelling the relative risk** can easily be fitted by many programs (not SPSS).

Logistic regression in STATA:

xi: logit obese age i.sex

or

xi: glm obese age i.sex, fam(bin) link(logit)

Relative risk model:

xi: glm obese age i.sex, fam(bin) link(id)

The **link** is **log** instead of **logit**

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10

Other methods to analysis of binary response data

Risk difference models

Logistic regression model focus on the Odds Ratios

This is the correct thing to do in **case-control** studies.

In **follow-up studies** **Risk Difference** is often the appropriate measure of association, (community effect).

I.e. a model like this might be more relevant:

$$\Pr(\text{event}) = p_0 + RD_1 + RD_2 + RD_3$$

$$\Pr(\text{event given the covariates}) = \alpha + \sum_{i=1}^p (\beta_i \cdot x_i)$$

That is linear on **probability scale**

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11

Other methods to analysis of binary response data

Risk difference models

$$\Pr(\text{event given the covariates}) = \alpha + \sum_{i=1}^p (\beta_i \cdot x_i)$$

Such a model **modelling the risk difference** can easily be fitted by many programs (not SPSS).

Logistic regression in STATA:

xi: logit obese age i.sex

or

xi: glm obese age i.sex, fam(bin) link(logit)

Risk difference model:

xi: glm obese age i.sex, fam(bin) link(id)

The **link** is **identity** instead of **logit**

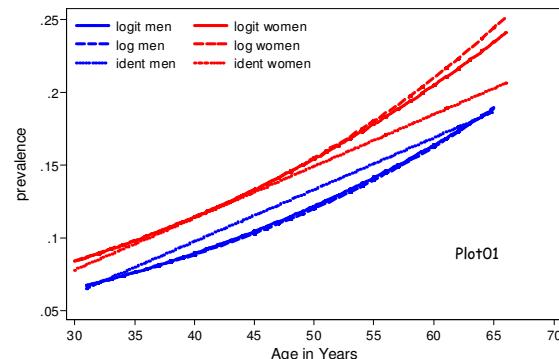
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12

Other methods to analysis of binary response data

Three different links for *Obese* "=" *sex* "+" *age*



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13

Other methods to analysis of binary response data Problems

$$\Pr(\text{event}) = p_0 \times RR_1 \times RR_2 \times RR_3$$

As the relative risk can be **larger** than one
the product might be **larger** than one !

$$\Pr(\text{event}) = p_0 + RD_1 + RD_2 + RD_3$$

The sum might **negative** and be **larger than one** !

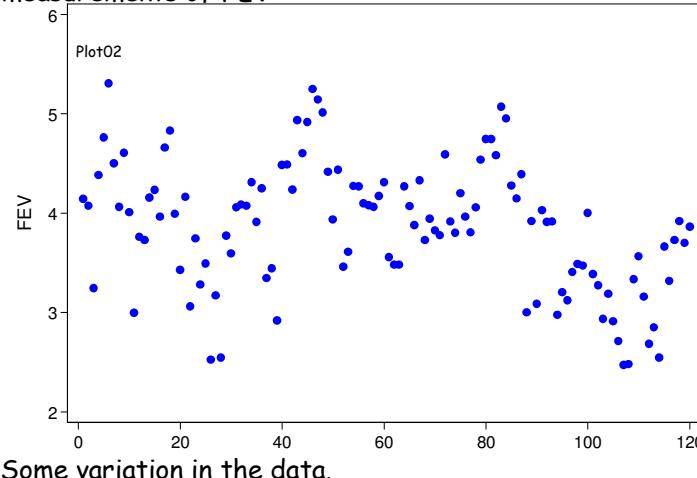
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14

Clustered data / data with several random components

120 measurements of FEV:



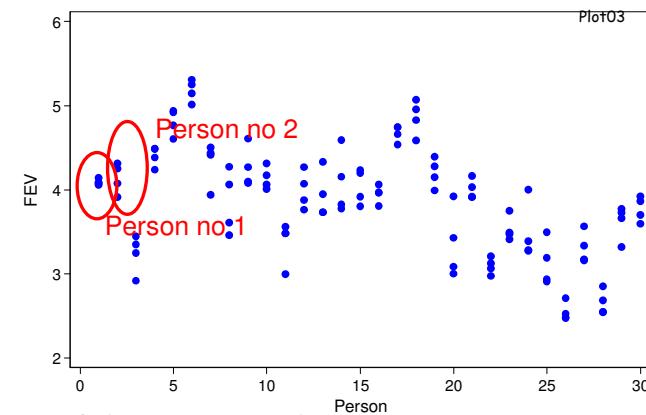
Some variation in the data.

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15

Clustered data / data with several random components But it is on only 30 persons:



Some of the variation is due to **variation between persons** and some within person.

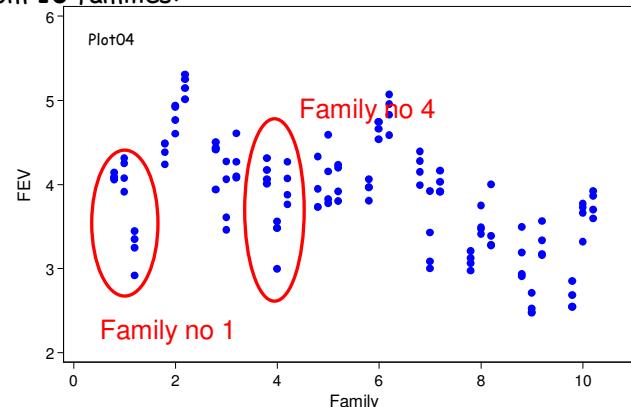
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16

Clustered data / data with several random components

From 10 families:



Some of the variation between persons is due to variation between families and some within family.

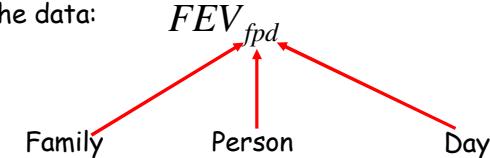
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17

Clustered data / data with several random components

Structure of the data:



Three sources of random variation:

Variation between families

Variation between persons (variation within family)

Variation between days (variation within person)

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18

Clustered data / data with several random components

Factors of interest:

household	Income	Constant within family
Urbanization		Constant within family
Age		Constant within person; varies within family
Sex		Constant within person; varies within family
Grass pollen		Constant within day; varies within person

A model:

$$FEV = \beta_0 + \beta_I \cdot I + \beta_U \cdot U + \beta_A \cdot A + \beta_S \cdot S + \beta_G \cdot G + \text{random variation}$$

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19

Clustered data / data with several random components

$$FEV = \beta_0 + \beta_I \cdot I + \beta_U \cdot U + \beta_A \cdot A + \beta_S \cdot S + \beta_G \cdot G + \text{random variation}$$

If the three levels/sources of random variation are not taken into account :

- The precision of the β_I and β_U are highly overestimated
- The precision of the β_A and β_S are overestimated
- The estimates of the β_I and β_U will be biased if the not all families are represented by the same number of persons and each person is measured the same number of times.
- The estimates of the β_A and β_S will be biased if the not all persons are measured the same number of times.

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20

Clustered data / data with several random components

$$FEV = \beta_0 + \beta_I \cdot I + \beta_U \cdot U + \beta_A \cdot A + \beta_S \cdot S + \beta_G \cdot G + F_f + P_{fp} + E_{fpd}$$

variance

F_f	: Random family contribution	σ_F^2
P_{fp}	: Random person contribution	σ_P^2
E_{fpd}	: Random day contribution	σ_E^2

$$\text{var}(FEV_{fpd}) = \sigma_F^2 + \sigma_P^2 + \sigma_E^2$$

Variance components

Assumed to be normal distributed

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21

Clustered data / data with several random components

Systematic part

$$FEV = \beta_0 + \beta_I \cdot I + \beta_U \cdot U + \beta_A \cdot A + \beta_S \cdot S + \beta_G \cdot G + F_f + P_{fp}$$

Random part

$\beta_0, \beta_I, \beta_U, \beta_A, \beta_S$ and β_G Quantify the **systematic** variation

σ_F^2, σ_P^2 and σ_E^2 Quantify the **random** variation

This is a:

- **Variance component model**
- **Mixed model** (both systematic and random variation)
- **Multilevel model**

The theory behind and the understanding of such models is well **established!!!**

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22

Clustered data / data with several random components

Dichotomous outcome

A different outcome:

$$H_{fpd} = \begin{cases} 1 & \text{if the person has hayfewer} \\ 0 & \text{else} \end{cases}$$

A statistical model:

Systematic part

$$\text{logit}(H_{fpd} = 1) = \beta_0 + \beta_I \cdot I + \beta_U \cdot U + \beta_A \cdot A + \beta_S \cdot S + \beta_G \cdot G + F_f + P_{fp}$$

Random part

This is not needed due to the binomial error

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23

Clustered data / data with several random components

Dichotomous outcome

$$\text{logit}(H_{fpd} = 1) = \beta_0 + \beta_I \cdot I + \beta_U \cdot U + \beta_A \cdot A + \beta_S \cdot S + \beta_G \cdot G + F_f + P_{fp}$$

That is, an ordinary logistic regression + **random components**.

- **A generalized linear mixed model**
- **A multilevel model for dichotomous outcome**

Comments 1:

- It is **important** to include the **relevant random components** in the model.
- 'Multilevel models' is **essential** in medical/epidemiological research.

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24

Clustered data / data with several random components Dichotomous outcome

Comments 2:

- The theory and insight into the models for non-normal data are **not yet fully developed**.
- The main problem being that it is not known how to obtain **valid (unbiased) estimates**.
- Several software programs **falsey claim** to estimate the models. (SAS, STATA, SUDAAN, NLwin)
- The programs/algorithms are not able give 'the correct' estimates.

Advice:

At the moment, do not trust results based on multilevel models.

Wait and see, the statisticians **might** solve the problems.

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25

Clustered data / data with one random components Dichotomous outcome

If the models only involves **one random components**, e.g. **variation between families** or between GP's,

then methods exists which can **adjust the standards errors**.

Remember that if the **data contains clusters**, then the precision of the estimates overestimated, that is the reported **standard errors is too small**.

So called **robust methods** or **sandwich estimates** of the standard errors will (try) adjust for this problem.

Only a **few** programs have this option - STATA does!

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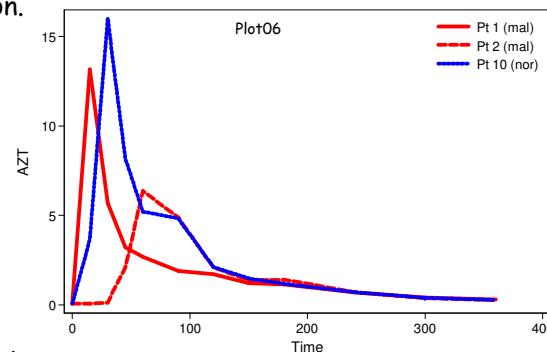
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26

Nonlinear regression models

Concentration in the blood of zidovudine (AZT) after administration of the drug.

One person with normal fat absorption and two with malabsorption.



Clearly non linear.

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27

Nonlinear regression models

The is **no way** that the above data can be described by a linear regression.

Furthermore **pharmaco kinetic** theory specify a simple model for the **expected concentration** as a function of time:

$$\text{concentration}(t) = d \frac{k_A}{k_A - k_E} (\exp(-k_E \cdot t) - \exp(-k_A \cdot t))$$

Where:

d : dose (per kg bodyweight)

k_A : absorptionsrate

k_E : eliminationsrate

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28

