

## Solution to exercise 4-3

### Background

It is a common belief that if the first two children in a family are of the same sex then the next child will most likely be of the same sex. All though scientific results are not completely consistent<sup>1</sup>, most studies favor independence between the sex of the children in a family. Here we revisit the hypothesis using data from a random sample of the first and second born in 1000 Danish families.

### Statistical methods

The chance of getting a boy the second time is analyzed using a binomial model and estimates with 95% confidence intervals (CI) are presented.

### Results

1. *Consider the chance of getting a boy the second time. Estimate this if the older sib was a boy and if she was a girl.*

The chances of getting a boy the second time is 52.7% (95% CI: 48.2%-57.2%) if the first child is a girl and 46.5% (42.1%-51.0%) if the first child is a boy.

2. *Estimate the risk difference comparing the chance of a boy the second time if the older sib was a boy to if the older sib was a girl. Test the hypothesis of no association between the sex of the first and the second child. Write a short conclusion including a discussion of the validity of the assumptions behind the calculations.*

The chance of getting a boy the second time is 6.2% (0.0%-12.4%) higher if the first child is a girl as compared to when the first child is a boy. The data suggest that the chance of getting a boy the second time may depend on the sex of the first born child ( $p=0.050$ ).

3. *Estimate the proportion of boys among the oldest and youngest sibs.*

The chances of getting a boy the first time is 50.3% (47.2%-53.4%) and the second time 49.6% (46.5%-52.7%).

4. *Estimate the difference in the proportion of boys among the first and second child. Test the hypothesis of no difference. Write a short conclusion including a discussion of the validity of the assumptions behind the calculations.*

The chance of getting a boy the second time is 0.7% (-3.9%-5.3%) lower as compared to chance of getting a boy the first time. The chance of getting a boy the first and second time is not statistically significant different ( $p=0.79$ ).

5. Discuss the similarities and differences between what you looked at in 2 and 4.

The analysis performed in question 2 examines the hypothesis of independence between the sex of the first and second child, whereas the analysis performed in question 4 examines the hypothesis of the same chance of having the a boy the first and second time.

### **Conclusion**

The chance of getting a boy the second time is 6.2% (12.4%-0.0%) higher if the first child is a girl as compared to when the first child is a boy. Although the data suggest that the chance of getting a boy the second time may depend on the sex of the first born child ( $p=0.050$ ), the finding is in opposite direction as the hypothesis of same sex in the family.

### **References**

1. [http://en.wikipedia.org/wiki/Human\\_sex\\_ratio#Natural\\_factors](http://en.wikipedia.org/wiki/Human_sex_ratio#Natural_factors)

## Do file

\*\*\*\*\*

\* Solution to Exercise 4-3.

\*\*\*\*\*

```
cd "D:\Teaching\BasicBiostat\Exercises"
```

```
capture log close
```

```
log using "solution4-3.log",text replace
```

```
use siblings.dta, clear
```

```
codebook sex*
```

```
tabu sex1 sex2
```

\* Many Stata commands expect 0/1 variables.

```
generate boy1=2-sex1 if sex1<.
```

```
generate boy2=2-sex2 if sex2<.
```

```
label define NoYes 0 "No" 1 "Yes"
```

```
label val boy1 NoYes
```

```
label val boy2 NoYes
```

```
tabu boy1 sex1
```

```
tabu boy2 sex2
```

\* Q1.

\* We use the tabulate command to get familiar with the counts,

\* and the ci command to computed the chance of getting a boy the

\* second time with 95% CI. The sex of the second child (boy2) is

\* the outcome and the sex of the first child (boy1) is the exposure.

```
tabu boy1 boy2,row
```

```
ci prop boy2 if boy1==0
```

```
ci prop boy2 if boy1==1
```

\* Q2

\* Risk difference, and hypothesis of independence.

\* Note:

\* - Since the sex of the second child is the outcome we place

\* boy2 first in the cs command and then the exposure boy1.

\* - The cs command does not use the labels coded in boy1 and boy2,

\* but rather the general terminology from cohort studies;

\* exposed (boy1=1, i.e. first born boy)

\* unexposed (boy1=0, i.e. first born girl)

\* cases (boy2=1, i.e. second born boy)

\* noncases (boy2=0, i.e. second born girl).

\* - The command prtest and tabu also performs the independence test

\* as in the cs command.

```
cs boy2 boy1
```

```
prtest boy2 , by(boy1)
```

```
tabu boy1 boy2 , row chi2
```

\* Q3

- \* Note: here the sex of the first (boy1) and second child (boy2) is the outcome.

ci prop boy1

ci prop boy2

\* Q4

- \* We can't at this point assume that the sex of the two siblings are independent, so we should consider the data as paired data.

\* Note:

- \* - The order of variable in mcc arbitrary, however we need in the output to keep track of who is placed where in the output table. Here we use the 2. variable first and then the 1. variable.
  - \* - The mcc command does not use the labels coded in boy1 and boy2, but rather the general terminology from matched case-control studies;
  - \* Cases (outcome of the second child, i.e. boy2)
    - \* exposed (boy2=1, i.e. second born boy)
    - \* unexposed (boy2=0, i.e. second born girl)
  - \* Control (outcome of the first child, i.e. boy1)
    - \* cases (boy1=1, i.e. first born boy)
    - \* noncases (boy1=0, i.e. first born girl).
  - \* - Be careful: prtest makes a unpaired comparison, i.e. prtest boy1=boy2
- mcc boy2 boy1

\* Q5

- \* Same probability of the sex is not the same as independence.

log close